Macro II: Stochastic Processes II

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Introduction

- ► Today: study a variety of stochastic processes that show up in macroeconomics.
- ► Then, discuss detrending data.

Stochastic Processes

- Conditional expectations and linear projections
- ▶ White noise
- ► AR(1)
- ► MA(1)
- ► ARMA(p,q)
- ▶ Detrending data

What is a Stochastic Process?

- Stochastic process is an infinite sequence of random variables $\{X_t\}_{t=-\infty}^{\infty}$
- ightharpoonup j'th autocovariance = $\gamma_j = C(X_t, X_{t-j})$
- ▶ Strict stationarity: distribution of $(X_t, X_{t+j_1}, X_{t+j_2}, ... X_{t+j_n},)$ does not depend on t
- ► Covariance stationarity: \bar{X}_t and $C(X_t, X_{t-j})$ do not depend on t

Defining a Conditional Density

- ▶ Work with random vector $\underline{x} = (X, Y) \sim F(x, y)$.
 - X and Y are random variables
 - x and y are realizations of the random variables
 - ightharpoonup F(x,y) is joint cumulative distribution
 - f(x, y) is joint density function

Conditional Variables and Independence

- Conditional probability
 - ▶ when $Pr(\underline{x} \in B) > 0$,

$$\Pr(\underline{x} \in A | \underline{x} \in B) = \Pr(A|B) = \frac{\Pr(A \cap B)}{\Pr(B)}.$$

- ▶ Conditional distribution F(y|x) (handles Pr(B) = 0)
 - ▶ Marginal distribution: $F_X(x) = \Pr(X \le x)$
 - ▶ F(y|x) is $Pr(Y \le y)$ conditional on $X \le x$

Independence: The random variables X and Y are independent if

$$F(x,y) = F_X(x) F_Y(y)$$

► If X and Y are independent, then

$$F(y|x) = F_Y(y)$$

and

$$F(x|y) = F_X(x)$$

- i.i.d means independent and identically distributed
- ► Conditional (mathematical, rational) expectation

$$E(Y|x) = \int_{-\infty}^{\infty} y dF(y|x) = \int_{-\infty}^{\infty} y f(y|x) dy.$$

Conditional Expectations as OLS

 Conditional expectation: function that minimizes the mean squared forecast error

$$E(Y|\underline{X}) = \arg\min_{\{f(\cdot)\}} E([Y - f(\underline{X})]^2)$$

where X is a vector

▶ Best linear predictor or linear projection: linear function that minimizes the mean squared forecast error

$$\hat{E}(Y|\underline{X}) = \arg\min_{\{\text{ linear } f(\cdot)\}} E([Y - f(\underline{X})]^2)$$

- $\hat{E}(Y|X) = E(Y|X)$ when E(Y|X) is linear
- ▶ OLS estimate since OLS gives best linear unbiased estimates

White noise

$$\{\varepsilon_t\}_{t=-\infty}^{\infty}$$

- $ightharpoonup E(\varepsilon_t) = 0, \ \forall t$
- $V(\varepsilon_t) = \sigma_{\varepsilon}^2, \ \forall t$

First-order autoregressive (AR(1)) process

$$x_t = \alpha + \phi x_{t-1} + \varepsilon_t$$

- lacksquare ε_t is white noise and $|\phi| < 1$ as required by stationarity
- ▶ By recursive substitution under stationarity

$$x_{t} = \alpha + \varepsilon_{t} + \phi \left[\alpha + \phi x_{t-2} + \varepsilon_{t-1} \right]$$
$$= \frac{\alpha}{1 - \phi} + \sum_{j=0}^{\infty} \phi^{j} \varepsilon_{t-j}.$$

$$E(x_t) = \alpha/(1-\phi)$$

Moments

Facts

$$V(aX + bY) = a^{2}V(X) + b^{2}V(Y) + 2abC(X, Y)$$

$$C(aX + bY, cX + dY) = acV(X) + bdV(Y) + (ad + bc)C(X, Y)$$

 \triangleright Since the value of x_t can be expressed as

$$x_t = \frac{\alpha}{1 - \phi} + \sum_{i=0}^{\infty} \phi^i \varepsilon_{t-i},$$

ightharpoonup The variance of x_t is

$$V(x_t) = \sum_{i=0}^{\infty} (\phi^i)^2 \sigma_{\varepsilon}^2 = \frac{\sigma_{\varepsilon}^2}{1 - \phi^2}.$$

Covariances

Covariance

$$C(x_{t}, x_{t-1}) = C(\alpha + \phi x_{t-1} + \varepsilon_{t}, x_{t-1})$$

$$= 0 + \phi V(X) + 0 = \phi \frac{\sigma_{\varepsilon}^{2}}{1 - \phi^{2}},$$

$$C(x_{t}, x_{t-k}) = C\left(\phi^{k} x_{t-k} + \sum_{j=0}^{k-1} \phi^{j} \varepsilon_{t-j}, x_{t-k}\right)$$

$$= \phi^{k} \frac{\sigma_{\varepsilon}^{2}}{1 - \phi^{2}} = \phi^{k} V(x_{t}).$$

▶ Expectation: If $\{\varepsilon_t\}$ is i.i.d. and $\alpha = 0$, $E(x_t|x_{t-k}) = \phi^k x_{t-k}$

AR(p)

▶ Autoregressive function of p lagged x's

$$\mathbf{x}_t = \alpha + \phi_1 \mathbf{x}_{t-1} + \phi_2 \mathbf{x}_{t-2} + \dots + \phi_p \mathbf{x}_{t-p} + \varepsilon_t$$

▶ Defining $x_{t-j} = L^j x_t$, we can rewrite an AR(p) process as

$$(1 - \phi_1 L - \phi_2 L^2 - \dots - \phi_p L^p) x_t = \alpha + \varepsilon_t$$

Stationarity condition: The roots of

$$1 - \phi_1 z - \phi_2 z^2 - \dots - \phi_p z^p = 0$$

lie outside the unit circle (|z| > 1 when real)

AR(p)

- ▶ For AR(1), ϕ < 1 yields stationarity
- ▶ If stationarity holds, we can rewrite x_t as a function of infinitely lagged $\varepsilon's$

First-order moving average (MA(1)) process

$$x_t = \alpha + \varepsilon_t + \theta \varepsilon_{t-1}$$

- \triangleright ε_t is white noise
- \triangleright $E(x_t) = \alpha$
- $V(x_t) = (1 + \theta^2) \sigma_{\varepsilon}^2$
- $C(x_t, x_{t-1}) = C(\varepsilon_t + \theta \varepsilon_{t-1}, \varepsilon_{t-1} + \theta \varepsilon_{t-2}) = \theta \sigma_{\varepsilon}^2$
- $C(x_t, x_{t-k}) = C(\varepsilon_t + \theta \varepsilon_{t-1}, \varepsilon_{t-k} + \theta \varepsilon_{t-k-1}) = 0, \ k > 1$

MA cont'd

Rewrite with lag operator as

$$x_t - \alpha = (1 + \theta L)\varepsilon_t$$

▶ When the root of

$$1 + \theta z = 0$$

lie outside unit circle (when $|\theta| < 1$) x_t is said to be invertible

$$\varepsilon_t = \frac{(x_t - \alpha)}{(1 + \theta L)}$$
$$= \frac{-\alpha}{1 + \theta} + \sum_{j=0}^{\infty} (-\theta)^j x_{t-j},$$

Express residual as infinite recursion of lagged x/s

MA(q)

 \triangleright x_t is a function of q lagged residuals

$$x_t = \alpha + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} + \dots + \theta_q \varepsilon_{t-q}$$

Rewriting with lag operator yields

$$(1 + \theta_1 L + \theta_2 L^2 + \dots + \theta_q L^q) \varepsilon_t = x_t - \alpha$$

Invertibility condition is that the roots of

$$1 + \theta_1 L + \theta_2 L^2 + \dots + \theta_q L^q = 0$$

lie outside the unit circle (|z| > 1 when real)

If the invertibility condition holds, we can write the ε_t as an infinite function of lagged x's

ARMA process

$$\begin{aligned} x_t &= \alpha + \phi_1 x_{t-1} + \phi_2 x_{t-2} + \ldots + \phi_p x_{t-p} \\ &+ \varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} + \ldots + \theta_q \varepsilon_{t-q}. \end{aligned}$$

- Stationarity condition
 - Depends entirely on autoregressive coefficients
 - The roots of

$$1 - \phi_1 z - \phi_2 z^2 - \dots - \phi_p z^p = 0,$$

must lie outside the unit circle ($|\cdot| > 1$ when real)

EX

► Example: For AR(1)

$$1 - \phi_1 z = 0$$

implying

$$z = \frac{1}{\phi_1}$$

need

$$|z|=|\phi_1^{-1}|>1$$
 requiring $|\phi_1|<1$

Invertibility Condition

► The roots of

$$1 - \theta_1 z - \theta_2 z^2 - \dots - \theta_p z^p = 0,$$

must lie on or outside the unit circle ($|\cdot| > 1$ when real)

► Example: For MA(1)

$$1 - \theta_1 z = 0$$

implying

$$z=rac{1}{ heta_1}$$

need

$$|z| = |\theta_1^{-1}| \ge 1$$
 requiring $|\theta_1| \le 1$,

where unity is included as a limit

Stationarity and invertibility

- Stationarity and invertibility imply
 - ▶ if ε_t is i.i.d., then ε_t is the innovation to x_t

$$\varepsilon_t = x_t - E(x_t | x_{t-1}, x_{t-2},).$$

knowledge of the entire sequences $\{\varepsilon_{t-j}\}_{j=0}^{\infty}$ and $\{x_{t-j}\}_{j=0}^{\infty}$ is equivalent

Detrending

- ► Most of our (business cycle) models have nothing to say about trends in the data.
- ▶ i.e., these models generally don't explain growth.
- Need to detrend to get an appropriate data series.
- $Y_t = X_t + z_t$
 - X_t is stationary
 - $ightharpoonup z_t$ is a trend
- Trend stationary
 - \triangleright z_t is deterministic
 - ightharpoonup example: $z_t = \alpha t$

Difference Stationary

- ► Difference stationary
 - $ightharpoonup z_t$ is a random walk with $\{\varepsilon_t\}$ a white noise process

$$z_{t} = z_{0} + \sum_{j=1}^{t} \varepsilon_{t-j}$$

$$z_{t-1} = z_{0} + \sum_{j=1}^{t-1} \varepsilon_{t-j}$$

$$z_{t} - z_{t-1} = \varepsilon_{t}$$

Three approaches

- Linear detrending: Regress data on time and take residuals
- Use Hodrick-Prescott filter to separate data into a trend component and residuals and take residuals
- First difference the data

HP Filter

- Most common approach: HP Filter.
- Idea: isolate low-frequency trends from high frequency cycles.
- ▶ Let $\{y_t\}_{t=1}^{\infty}$ be a given series, where $y_t = x_t + z_t$ as before.
- \triangleright x_t is the trend component, z_t is cyclical.
- Let λ be a parameter to be specified later, and consider the problem

$$\min_{x_1, x_2, \dots, x_T} \sum_{t=1}^{T} (y_t - x_t)^2 + \lambda \sum_{t=2}^{T-1} [(x_{t+1} - x_t) - (x_t - x_{t-1})]^2$$
(1)

- ▶ What is going on here?
 - We are minimizing the cyclical component (first part), by moving the trend closer to the data.
 - ▶ But we are getting penalized (λ) for making the trend too closely reflect the data.

Next Time

- Discuss expectations in linear difference equations.
- ▶ Please turn your homework in by this evening.
- See my webpage for new homework (may not be up by class).